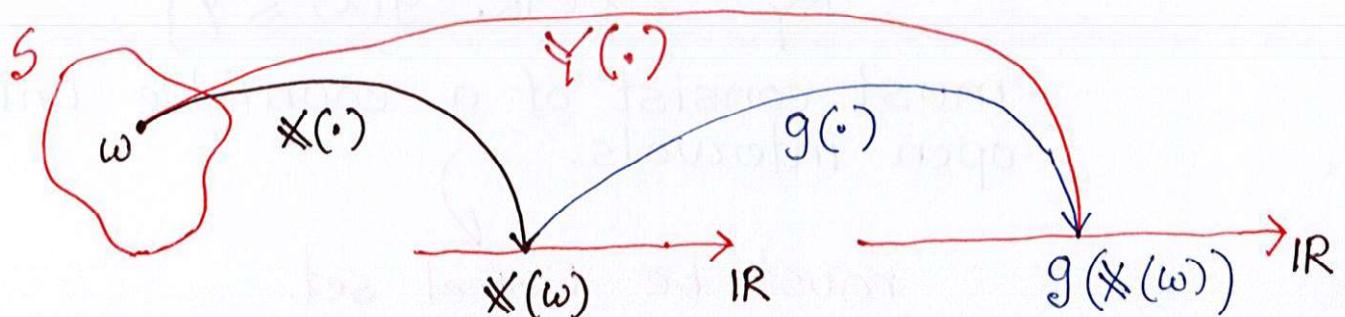


## Functions of Random Variables

Let's assume that  $\mathbb{X}$  is a random variable on the probability space  $(S, \mathcal{F}, P)$ .

We can use this random variable  $\mathbb{X}$  and define a function  $\mathbb{Y}$  as follows:

$$\mathbb{Y} = g(\mathbb{X}), \text{ where } g: \mathbb{R} \rightarrow \mathbb{R}$$



As we see, the mapping of  $w$  is done to  $\mathbb{R}$  twice when we define a function of random variable.

A composite function structure is evident from the definition

$$\mathbb{Y}(w) = g(\mathbb{X}(w)), \text{ that is}$$

$$\mathbb{Y}: S \rightarrow \mathbb{R}$$

However, the question is

Is  $\mathbb{Y}(\cdot)$  a random variable?

To assess it, let's recall the definition:

$\mathbb{Y}: S \rightarrow \mathbb{R}$  is a random variable if

$$\mathbb{Y}^{-1}(A) = \{\omega \in S \mid \mathbb{Y}(\omega) \in A\} \in \mathcal{F},$$
$$\forall A \in \mathcal{B}(\mathbb{R})$$

where,  $\mathcal{F}$  is the event space of  $(S, \mathcal{F}, P)$

For  $Y = g(X)$  to be measurable (that is random variable)  $g(\cdot)$  must satisfy the following properties

1. The domain of  $g(\cdot)$  must contain the range space of  $X$
2. For each  $y \in \mathbb{R}$ , the set  $R_y$  defined as  $R_y = \{x \in \mathbb{R}; g(x) \leq y\}$  must consist of a countable union of open intervals.  
must be Borel set
3. The events  $\{g(X) = \pm\infty\}$  must have probability zero.

So, any function  $g(\cdot)$  that satisfies these 3 properties is known as Baire function.

For such  $g(\cdot)$ , we can say that

$$Y = g(X)$$

is a valid random variable.

Interestingly,

All functions we typically encounter in engineering applications are Baire functions.

Example :  $Y = g(X) = X^2 \rightarrow \text{fn}^c$

The transformation considers  $g(x) = x^2 = Y$   
As any specific  $y$  is the square of  $x$ ,  $y$   
will always be non-negative. So,

case  $y < 0$  :  $F_Y(y) = 0, y < 0$  [negative value  
of  $y$  is impossible]

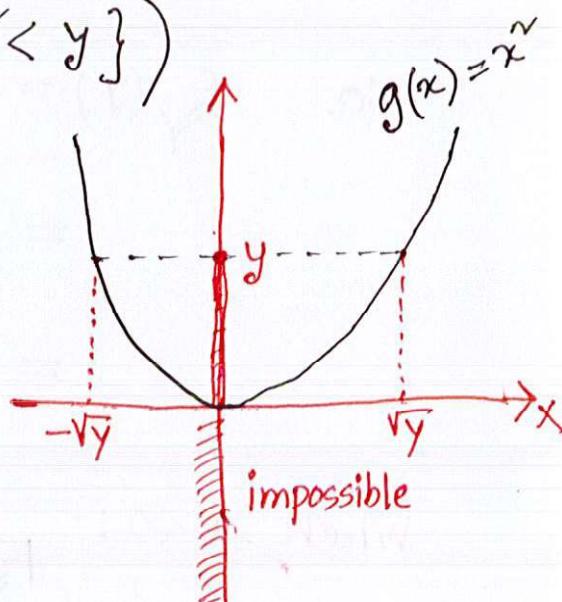
For the case  $y > 0$  : Positive  $y$

$$F_Y(y) = P(\{Y \leq y\}) = P(\{X^2 \leq y\})$$

$$= P(\{-\sqrt{y} \leq X \leq +\sqrt{y}\})$$

$$= P(\{-\sqrt{y} < X \leq +\sqrt{y}\})$$

or, cdf is continuous  
for continuous RV  
case zero probability  
event can be written  
without the equal sign



$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

So, the CDF is

$$F_Y(y) = [F_X(\sqrt{y}) - F_X(-\sqrt{y})] \cdot \mathbb{1}_{(y, \infty)}$$

what about  $y=0$ ?  $y$  becomes zero only at one point  $x$ , that is,  $x=0$ . Now, if the cdf is continuous at zero, then the probability that  $x=0$  is equal to zero.

So, we can ignore the case  $(y, \infty)$

$$f_Y(y) = \frac{d}{dy} F_Y(y)$$

Examples:  $Y = g(X) = aX + b$ ,  $a, b \in \mathbb{R}$

Find  $f_Y(y)$ . Two cases  $a > 0$  and  $a < 0$

When  $a > 0$ :  $F_Y(y) = P\{Y \leq y\} = P\{aX + b \leq y\}$

$$= P\left\{X \leq \frac{y-b}{a}\right\}$$

$$= F_X\left(\frac{y-b}{a}\right)$$

so,  $f_Y(y) = \frac{d}{dy} F_Y(y)$

$$= \frac{d}{dy} F_X\left(\frac{y-b}{a}\right) = f_X\left(\frac{y-b}{a}\right) \cdot \frac{d}{dy}\left(\frac{y-b}{a}\right)$$

$$= f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} = \frac{1}{a} f_X\left(\frac{y-b}{a}\right)$$

When  $a < 0$ :  $F_Y(y) = P\{Y \leq y\}$

$$= P\{aX + b \leq y\} = P\left\{X \leq \frac{y-b}{a}\right\}$$

As  $a < 0$ , it is negative, so,

$$= P\left\{X > \frac{y-b}{a}\right\}$$

because  $a < 0$ , the inequality changes

$$= 1 - P\left\{X \leq \frac{y-b}{a}\right\}$$

$$= 1 - F_X\left(\frac{y-b}{a}\right)$$

so,  $f_Y(y) = \frac{d}{dy} \left[ 1 - F_X\left(\frac{y-b}{a}\right) \right]$

$$\Rightarrow f_Y(y) = -f_X\left(\frac{y-b}{a}\right) \cdot \frac{1}{a} = -\frac{1}{a} f_X\left(\frac{y-b}{a}\right)$$

By combining  $a < 0$ ,  $a > 0$ , we can write

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y-b}{a}\right)$$

$$\text{So, } f_Y(y) = \frac{dF_Y(y)}{dy}$$

$$\Rightarrow f_Y(y) = \frac{d}{dy} [F_X(\sqrt{y}) - F_X(-\sqrt{y})] \cdot 1_{(y, \infty)}^{(y)}$$

$$\begin{aligned}\Rightarrow f_Y(y) &= f_X(\sqrt{y}) \frac{d}{dy} (\sqrt{y}) - f_X(-\sqrt{y}) \frac{d}{dy} (-\sqrt{y}) \\ &= f_X(\sqrt{y}) \frac{1}{2\sqrt{y}} - f_X(-\sqrt{y}) \frac{-1}{2\sqrt{y}} \\ &= \frac{1}{2\sqrt{y}} [f_X(\sqrt{y}) + f_X(-\sqrt{y})] \\ &= \frac{1}{2\sqrt{y}} [f_X(\sqrt{y}) + f_X(-\sqrt{y})] \cdot 1_{(0, \infty)}^{(y)}\end{aligned}$$

## Function of random variables

The direct pdf method:

Suppose  $Y = g(X)$ , where  $g: \mathbb{R} \rightarrow \mathbb{R}$  such that  $g^{-1}(\cdot)$  exists. That is  $y = g(x)$

$$\Rightarrow x = g^{-1}(y)$$

$$\Rightarrow x(y) = g^{-1}(y)$$

It is also assumed that

$$\frac{dx}{dy} = \frac{d g^{-1}(y)}{dy} \text{ exists.}$$

$$\text{Then, } f_Y(y) = f_X(g^{-1}(y)) \cdot \left| \frac{d g^{-1}(y)}{dy} \right|$$

$$\text{where, } x(y) = g^{-1}(y)$$

**Example:** Given that  $X \sim U[0, 1]$  and let's assume that  $Y = g(X) = \sqrt{X}$ . So, find  $f_Y(y)$ .

Let's apply direct pdf method

According to the formula :

$$\text{Here, } y = g(x) = \sqrt{x}$$

$$\Rightarrow x = y^2$$

$$\Rightarrow x(y) = y^2$$

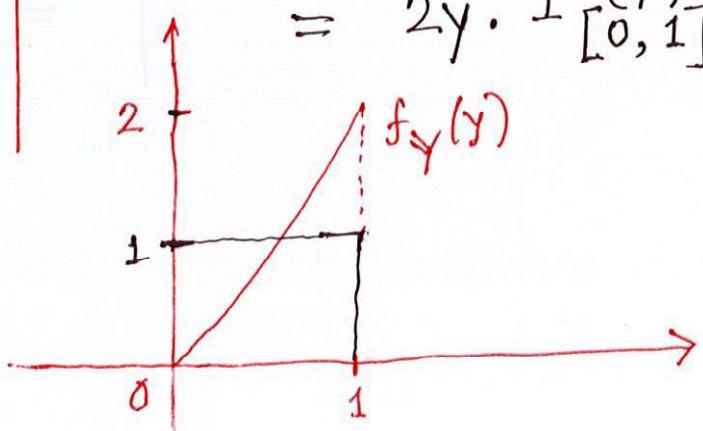
$$\text{so, } \frac{dx(y)}{dy} = 2y$$

$$f_Y(y) = f_X(x(y)) \left| \frac{dx(y)}{dy} \right|$$

$$= f_X(g^{-1}(y)) \left| \frac{d g^{-1}(y)}{dy} \right|$$

$$\text{so, } f_Y(y) = |2y| \cdot f(X(y))$$

$$= 2y \cdot 1_{[0,1]}(y)$$



Example: Consider  $\tilde{x}$  be a Gaussian Random variable with pdf

$$f_{\tilde{x}}(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

$N(\mu, \sigma^2)$  is the standard notation

mean  $\mu: 0$   
variance  $\sigma^2: 1$

for the Gaussian Random Variable with mean  $\mu$  and variance  $\sigma^2$ .

Consider the linear transformation  $y = ax + b$   
Let's find out  $f_y(y)$  using the direct pdf method.

$$f_y(y) = f_{\tilde{x}}(x(y)) \left| \frac{dx(y)}{dy} \right|$$

$$\begin{aligned} y &= ax + b \\ \Rightarrow x &= \frac{y-b}{a} \\ \Rightarrow \underline{x(y)} &= \frac{y-b}{a} \end{aligned}$$

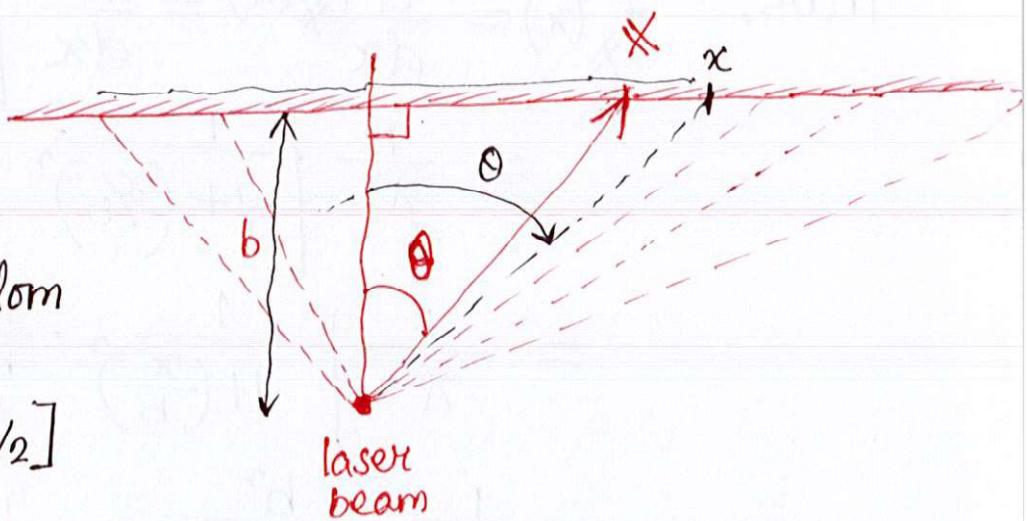
↳ rewriting  
considering  $x$  as  
a func of  $y$

$$\begin{aligned} \frac{dx(y)}{dy} &= \frac{d}{dy} \left( \frac{y-b}{a} \right) \\ &= \frac{1}{a} \end{aligned}$$

$$\begin{aligned} \therefore f &= f_{\tilde{x}}(x(y)) \left| \frac{1}{a} \right| \\ &= \frac{1}{\sqrt{2\pi}} \exp\left(\frac{-(y-b)^2}{2a^2}\right) \cdot \frac{1}{|a|} \\ &= \frac{1}{\sqrt{2\pi}|a|} \exp\left\{\frac{-(y-b)^2}{2a^2}\right\} \end{aligned}$$

**Example:** Suppose that a laser beam points at infinitely long wall, and the wall is at a distance  $b$ .

Assume that  $\theta$  is uniformly distributed random variable on the interval  $[-\pi/2, \pi/2]$



Let's say  $X$  is the coordinate along the wall where the laser hits. Find the pdf of  $X$

**Solution:** Let's find out  $F_X(x)$  first.

$$\begin{aligned}\tan \theta &= \frac{x}{b} \\ \Rightarrow \theta &= \tan^{-1} \frac{x}{b} \\ \Rightarrow x &= b \tan \theta \\ \Rightarrow x(\theta) &= b \tan \theta\end{aligned}$$

so,  $P(X \leq x) \xrightarrow{\text{random, changes for } \theta \text{ random}} = P(b \tan \theta \leq x) = P(\theta \leq \tan^{-1} \frac{x}{b}) = F_\theta(\tan^{-1} x/b)$

Now,  $F_X(x) = P\{X \leq x\}$

$$= F_\theta(\tan^{-1} x/b)$$

$\Rightarrow \theta$  values generating wall coordinate less than or equal to  $x$

————— total possible  $\theta$  values

$$= \frac{\tan^{-1}(x/b) + \pi/2}{\pi}$$

$$\text{So, } F_{\hat{x}}(x) = \frac{1}{\pi} \left[ \tan^{-1}(x/b) + \frac{\pi}{2} \right]$$

$$\begin{aligned} \text{Thus, } f_{\hat{x}}(x) &= \frac{d}{dx} F_{\hat{x}}(x) = \frac{d}{dx} \left[ \frac{1}{\pi} \left( \tan^{-1}(x/b) + \frac{\pi}{2} \right) \right] \\ &= \frac{1}{\pi} \left[ \frac{1}{1+(x/b)^2} \cdot \frac{d}{dx}(x/b) + 0 \right] \\ &= \frac{1}{\pi} \left[ \frac{1}{1+(\frac{x}{b})^2} \cdot \frac{1}{b} \right] = \frac{1}{\pi b} \left( \frac{1}{b^2+x^2} \right) \\ &= \frac{1}{\pi b} \cdot \frac{b^2}{b^2+x^2} \end{aligned}$$

Answer

**Question:** Assume that  $\hat{x}$  is a random variable uniformly distributed on  $(0, 1)$  and let  $Y = g(\hat{x})$ , where  $g(x) = -\lambda \ln(1-x)$

and  $\lambda > 0$

$$\text{Here, } f_{\hat{x}}(x) = 1_{[0, 1]}^{(x)}$$

According to direct method :  $f_Y(y) = f_{\hat{x}}(x(y)) \left| \frac{dx(y)}{dy} \right.$

$$\text{So, } y = g(x) = -\lambda \ln(1-x)$$

$$\Rightarrow -\frac{y}{\lambda} = \ln(1-x)$$

$$\Rightarrow 1-x = e^{-y/\lambda}$$

$$\Rightarrow x = 1 - e^{-y/\lambda}$$

$$\Rightarrow x(y) = 1 - e^{-y/\lambda}$$

$$\begin{aligned} \frac{d}{dy} x(y) &= \frac{d}{dy} (1 - e^{-y/\lambda}) \\ &= -e^{-y/\lambda} \cdot (-1) \\ &= e^{-y/\lambda} \cdot \frac{1}{\lambda} \end{aligned}$$

$$\text{So, } f_Y(y) = e^{-y/\lambda} \cdot \frac{1}{\lambda} \cdot 1_{[0, 1]}^{(y)}$$