

## Mean, Variance and Expectation

**Definition:** The mean or expected value of a random variable  $\mathbf{x}$  with pdf  $f_{\mathbf{x}}(x)$  is defined as:

$$E[\mathbf{x}] \triangleq \int_{-\infty}^{\infty} x f_{\mathbf{x}}(x) dx \quad \dots \dots \textcircled{1}$$

While the definition as in  $\textcircled{1}$  is for continuous RV, the definition can be extended for discrete RV as well.

For instance, pdf as  $\delta$ -function

If  $P(\{\mathbf{x} = x_k\}) = P_{\mathbf{x}}(x_k) = p_k$  over a discrete index set, then

$$f_{\mathbf{x}}(x) = \sum_k P_{\mathbf{x}}(x_k) \delta(x - x_k) = \sum_k p_k \delta(x - x_k)$$

and,  $E[\mathbf{x}] = \int_{-\infty}^{\infty} x f_{\mathbf{x}}(x) dx = \int_{-\infty}^{\infty} x \left( \sum_k p_k \delta(x - x_k) \right) dx$

$$= \sum_k p_k \int_{-\infty}^{\infty} x \delta(x - x_k) dx = \sum_k p_k x_k$$

$$= \sum_k P_{\mathbf{x}}(x_k) x_k \quad \text{Mean for discrete RV}$$

So, for a discrete RV  $\mathbf{x}$ , we have

$$E[\mathbf{x}] = \sum_k x_k P_{\mathbf{x}}(x_k)$$

## Conditional mean of RV

Let's consider  $\mathbf{X}$  be a RV on the probability space  $(S, \mathcal{F}, P)$  and  $M$  is an event in the event space.

Then, the conditional mean of  $\mathbf{X}$  conditioned on  $M$  is

$$E[\mathbf{X}|M] \triangleq \int_{-\infty}^{\infty} x f_{\mathbf{X}}(x|M) dx$$

However, if  $\mathbf{X}$  is a discrete RV, the conditional pmf  $P_{\mathbf{X}}(x_k|M) = P(\{\mathbf{X}=x_k\}|M)$ , then

$$\begin{aligned} E[\mathbf{X}|M] &= \int_{-\infty}^{\infty} x f_{\mathbf{X}}(x|M) dx = \int_{-\infty}^{\infty} x \left( \sum_k P_{\mathbf{X}}(x_k|M) \delta(x-x_k) \right) dx \\ &= \sum_k P_{\mathbf{X}}(x_k|M) \cdot \underbrace{\int_{-\infty}^{\infty} x \delta(x-x_k) dx}_{\text{dirac-}\delta\text{ property}} \\ &= \sum_k x_k P_{\mathbf{X}}(x_k|M) \end{aligned}$$

Mean: exponential

Let  $X$  be an exponentially distributed RV with pdf  $f_X(x) = \frac{1}{\mu} e^{-x/\mu} \mathbf{1}_{[0, \infty]}(x)$ ,  $\mu > 0$

What is  $E[X]$ ?

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^{\infty} \underbrace{x}_{u} \cdot \frac{1}{\mu} \underbrace{e^{-x/\mu}}_v dx$$

$= \int u v dx$ , where,  $u$  is a fn<sup>c</sup> of  $x$   
 $v$  is a fn<sup>c</sup> of  $x$

$$= \frac{1}{\mu} \left[ x \int e^{-x/\mu} dx - \int \frac{d}{dx} x \left( \int e^{-x/\mu} dx \right) dx \right]_0^{\infty}$$

$$= \frac{1}{\mu} \left[ -x \cdot e^{-x/\mu} \cdot \mu - \int -\mu e^{-x/\mu} dx \right]_0^{\infty}$$

$$= \left[ -x e^{-x/\mu} + \int e^{-x/\mu} \right]_0^{\infty}$$

$$= \left[ -x e^{-x/\mu} - \mu e^{-x/\mu} \right]_0^{\infty}$$

$$= -(-\mu e^{-0/\mu}) = \mu e^{-0} = \mu \text{ Mean}$$

We can use the same  $X$  and conditioned on  $M$ . What would be the expression for  $E[X|M]$ , where  $M = \{X > \mu\}$

According to expectation formula:

$$E[X|M] = E[X|\{X > \mu\}]$$

$$= \int_{-\infty}^{\infty} x f_X(x|\{X > \mu\}) dx$$

We can calculate  $f_{\hat{x}}(x / \underbrace{\{x > \mu\}}_M)$  by evaluating the cdf  $F_{\hat{x}}(x/M)$ .

Then, we can take derivative of  $F_{\hat{x}}(x/M)$  to obtain the cdf.

We can use the formula as follows:

$$F_{\hat{x}}(x/M) = P(\{x \leq x\}/M)$$

$$= \frac{P(\{x \leq x\} \cap M)}{P(M)} = \frac{P(A/\{x \leq x\}) F_{\hat{x}}(x)}{P(M)}$$

If,

$$M = \{x > \mu\}$$

$$= \frac{P(\{x > \mu\} / \{x \leq x\}) F_{\hat{x}}(x)}{P(\{x > \mu\})}$$

Question is,

How can you calculate  $F_{\hat{x}}(x)$  from the given  $f_{\hat{x}}(x)$ ?

 we need this for  $F_{\hat{x}}(x)$  calculation

$$F_{\hat{x}}(x) = \int f_{\hat{x}}(x) dx$$

$$F_{\hat{x}}(x/M) = \int f_{\hat{x}}(x/M) dx \quad \left| \begin{array}{l} P(\{a < x \leq b\}/M) \\ = F_{\hat{x}}(b/M) - F_{\hat{x}}(a/M) \end{array} \right.$$

We can write :

$$P(\{a < x \leq b\}/M) = \int_a^b f_{\hat{x}}(x/M) dx$$

so,  $E[X/\{X>\mu\}] = \mu + \mu = 2\mu$ . As we obtain  $E[X] \neq E[X|\{X>\mu\}]$ . Generally,

$$E[X] \neq E[X|M], M \in \mathcal{F}$$

田 RV as a fn<sup>c</sup> of another RV: Expectation

Suppose we have a RV  $X$  defined on  $(S, \mathcal{F}, P)$  and the pdf is  $f_X(x)$ .

Consider that a new random variable is defined as:  $Y = g(X)$ , where  $g: \mathbb{R} \rightarrow \mathbb{R}$

so, calculate

$$E[Y] = E[g(X)], \text{ where } g: \mathbb{R} \rightarrow \mathbb{R}$$

According to the formula, we can calculate as

$$E[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy$$

↳ we need evaluation of this.

we can use the below result to calculate  $E[Y]$ . So,

$$E[g(X)] \triangleq \int_{-\infty}^{\infty} g(x) f_X(x) dx$$

For a discrete random variable,  $f_X(x) = \sum_k p_X(x_k) \delta(x - x_k)$

$$E[g(X)] = \int_{-\infty}^{\infty} g(x) \left( \sum_k p_X(x_k) \delta(x - x_k) \right) dx$$

$$= \sum_k p_X(x_k) \int_{-\infty}^{\infty} g(x) \delta(x - x_k) dx = \sum_k g(x_k) p_X(x_k)$$

So, let's calculate the conditional mean  $E[\mathbb{X}/M]$  for a random variable with pdf

$$f_{\mathbb{X}}(x) = \frac{1}{\mu} \exp\left\{-\frac{x}{\mu}\right\} \cdot 1_{[0, \infty)}^{(x)}$$

$$\begin{aligned} E[\mathbb{X}/M] &= \int x f_{\mathbb{X}}(x/\mathbb{M}) dx \\ &= \int x f_{\mathbb{X}}(x/\{\mathbb{X} > \mu\}) dx \end{aligned} \quad \left| \begin{array}{l} \text{For} \\ M = \{\mathbb{X} > \mu\} \end{array} \right.$$

We can show that

$$\begin{aligned} f_{\mathbb{X}}(x/\{\mathbb{X} > \mu\}) &= \frac{\frac{1}{\mu} \exp\left\{-\frac{x}{\mu}\right\} \cdot 1_{(\mu, \infty)}^{(x)}}{e^{-\mu/\mu}} \\ &= \frac{1}{\mu} \exp\left\{-\frac{(x-\mu)}{\mu}\right\} 1_{(\mu, \infty)}^{(x)} \end{aligned}$$

$$\begin{aligned} \therefore E[\mathbb{X}/\{\mathbb{X} > \mu\}] &= \int_{-\infty}^{\infty} x f_{\mathbb{X}}(x/\{\mathbb{X} > \mu\}) dx \\ &= \int_{-\infty}^{\infty} x \cdot \frac{1}{\mu} \exp\left\{-\frac{(x-\mu)}{\mu}\right\} dx \end{aligned}$$

consider  $r = x - \mu$

so, we obtain,

$$\begin{aligned} &= \int_{-\infty}^{\infty} (r + \mu) \frac{1}{\mu} \exp\left\{-\frac{r}{\mu}\right\} dr \quad \left| \begin{array}{l} \Rightarrow x = r + \mu \\ \Rightarrow dx = dr \end{array} \right. \\ &= \int_{-\infty}^{\infty} r \cdot \frac{1}{\mu} e^{-\frac{r}{\mu}} dr + \int_{-\infty}^{\infty} \mu \cdot \frac{1}{\mu} e^{-\frac{r}{\mu}} dr \\ &= \frac{1}{\mu} \int_0^{\infty} r \cdot e^{-\frac{r}{\mu}} dr + \int_0^{\infty} \mu \cdot e^{-\frac{r}{\mu}} dr \\ &\quad \boxed{\left. [-\mu e^{-r/\mu}] \right|_0^{\infty} = +\mu \cdot e^0} \\ &\quad \boxed{0 = \mu \cdot 1} \\ &\quad \boxed{= \mu} \end{aligned}$$

From  $E[\mathbb{X}]$   
derivation

**Variance:** Given a random variable  $\mathbb{X}$ , the variance of  $\mathbb{X}$ , denoted as  $\text{Var}(\mathbb{X})$ , is as follows

$$\begin{aligned}\text{Var}(\mathbb{X}) &\triangleq E[(\mathbb{X} - \bar{\mathbb{X}})^2] \\ &= \int_{-\infty}^{\infty} (\mathbb{X} - \bar{\mathbb{X}})^2 f_{\mathbb{X}}(\mathbb{X}) d\mathbb{X}\end{aligned}\quad \left| \begin{array}{l} \text{Here,} \\ \bar{\mathbb{X}} = E[\mathbb{X}] \end{array} \right.$$

often  $\sigma_{\mathbb{X}}^2$  is used as the symbol for  $\text{Var}(\mathbb{X})$

$$\sigma_{\mathbb{X}}^2 = \text{Var}(\mathbb{X})$$

Positive square root of the variance of  $\mathbb{X}$  is called the standard deviation of  $\mathbb{X}$ .

$$\text{Std. Dev}(\mathbb{X}) = \sqrt{\sigma_{\mathbb{X}}^2} = \sigma_{\mathbb{X}} = \sqrt{\text{Var}(\mathbb{X})}$$

$$\begin{aligned}\text{Var}(\mathbb{X}) &= E[(\mathbb{X} - \bar{\mathbb{X}})^2] = E[\mathbb{X}^2 - 2\mathbb{X}\bar{\mathbb{X}} + \bar{\mathbb{X}}^2] \\ &= E[\mathbb{X}^2] - E[2\mathbb{X}\bar{\mathbb{X}}] + E[\bar{\mathbb{X}}^2] \\ &= E[\mathbb{X}^2] - 2\bar{\mathbb{X}}E[\mathbb{X}] + E[\bar{\mathbb{X}}^2] \\ &= E[\mathbb{X}^2] - 2\bar{\mathbb{X}}\bar{\mathbb{X}} + \bar{\mathbb{X}}^2 \\ &= E[\mathbb{X}^2] - \bar{\mathbb{X}}^2 = E[\mathbb{X}^2] - (E[\mathbb{X}])^2\end{aligned}$$

$$\text{So, } \text{Var}(\mathbb{X}) = E[\mathbb{X}^2] - (E[\mathbb{X}])^2$$

## 田 Linearity of expectation

consider that  $g_1(\mathbf{x})$  and  $g_2(\mathbf{x})$  are the functions of RV  $\mathbf{x}$ , and  $\alpha, \beta$  are the constants with  $\alpha, \beta \in \mathbb{R}$

$$\text{Then, } E[\alpha g_1(\mathbf{x}) + \beta g_2(\mathbf{x})] = \alpha E[g_1(\mathbf{x})] + g_2(\mathbf{x})\beta$$

**Example:** Consider a Gaussian RV  $\mathbf{x}$  with pdf

$$\mu \in \mathbb{R} \text{ and } \sigma > 0 \quad | \quad f_{\mathbf{x}}(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$$

Find  $E[\mathbf{x}]$  and  $\text{Var}(\mathbf{x})$

$$E[\mathbf{x}] = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} \cdot x \, dx$$

$$= \int_{-\infty}^{\infty} (\mu + r) \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\} dr \quad \begin{aligned} \text{Let } x - \mu &= r \\ \Rightarrow x &= \mu + r \\ \Rightarrow dx &= d\mu + dr \\ \Rightarrow dr &= dx \end{aligned}$$

$$= \int_{-\infty}^{\infty} (\mu + r) \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr$$

$$= \int_{-\infty}^{\infty} r \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr + \int_{-\infty}^{\infty} \mu \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr$$

expand it Part A

$$\text{Part A : } \frac{1}{\sqrt{2\pi}\sigma} \left[ \int_{-\infty}^0 r \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr + \int_0^{\infty} r \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr \right]$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \left[ - \int_0^{-\infty} r \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr + \int_0^{\infty} r \exp\left\{-\frac{r^2}{2\sigma^2}\right\} dr \right]$$

We can rewrite by doing a change of variable.

So,

$$\frac{1}{\sqrt{2\pi} 6} \int_{-\infty}^{\infty} (-r) \exp\left\{-\frac{r^2}{26^2}\right\} dr$$

consider  
 $(-r) = r$

$$= \frac{1}{\sqrt{2\pi} 6} \int_0^{\infty} (-r) \exp\left\{-\frac{(-r)^2}{26^2}\right\} dr$$

so,  $dr = -dr$

$$= \frac{1}{\sqrt{2\pi} 6} \int_0^{\infty} r \exp\left\{-\frac{r^2}{26^2}\right\} (-dr)$$

when  $-r = 0$   
 $\Rightarrow r = 0$   
 $-r = -\infty$   
 $\Rightarrow r = \infty$

after the change  
of variable process

$$= \frac{1}{\sqrt{2\pi} 6} \cdot (-1) \cdot \int_0^{\infty} r \exp\left\{-\frac{r^2}{26^2}\right\} dr$$

$$= -\frac{1}{\sqrt{2\pi} 6} \int_0^{\infty} r \exp\left\{-\frac{r^2}{26^2}\right\} dr$$

limits are changed

By plugging in these values, we obtain

$$\text{Part A} = \frac{1}{\sqrt{2\pi} 6} \left[ \int_0^{\infty} -r \exp\left\{-\frac{r^2}{26^2}\right\} dr + \int_0^{\infty} r \exp\left\{-\frac{r^2}{26^2}\right\} dr \right]$$

$$= 0$$

That is,

$$E[X] = \int_{-\infty}^{\infty} \mu \cdot \frac{1}{\sqrt{2\pi} 6} \exp\left\{-\frac{r^2}{26^2}\right\} dr$$

We can do change of variable by assuming

$$\frac{r^2}{26^2} = Z$$

$$\text{So, } z = \frac{x}{\sqrt{26}} \quad \Rightarrow \quad x^2 = 2z^2$$

$$\Rightarrow dz = \frac{1}{\sqrt{26}} dx$$

$$\Rightarrow dz \cdot \sqrt{26} = dx \quad \Rightarrow \quad \sqrt{26} dz = dx$$

Replacing the variable  $x$ -term by  $z$ , we obtain

$$\begin{aligned} E[x] &= \int_{-\infty}^{\infty} \mu \cdot \frac{1}{\sqrt{2\pi} \sigma} \exp\left\{-\left(\frac{x}{\sqrt{26}}\right)^2\right\} dx \\ &= \int_{-\infty}^{\infty} \mu \cdot \frac{1}{\sqrt{2\pi} \sigma} e^{-z^2} dz. \\ &= \int_{-\infty}^{\infty} \mu \cdot \frac{1}{\sqrt{\pi}} e^{-z^2} dz \\ &= \mu \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-z^2} dz = \mu \cdot \frac{2}{\sqrt{\pi}} \int_0^{\infty} e^{-z^2} dz \\ &= \mu \cdot 1 = \mu \end{aligned}$$

We know that  $\frac{2}{\sqrt{\pi}} \int_0^{\infty} e^{-z^2} dz = \lim_{z \rightarrow \infty} \int_0^z \frac{2}{\sqrt{\pi}} e^{-z^2} dz = 1$ .