

Conditional Probability

In many real-world scenarios, we cannot find the precise outcome of an experiment. Instead, we can just say that the event B occurs and the precise outcome ω_i is in the set B .

Overall, we just know that event B has occurred.

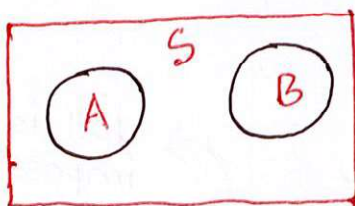
Assume that two events $A, B \in \mathcal{F}(S)$, and B has occurred.

↪ But we don't know the precise outcome.

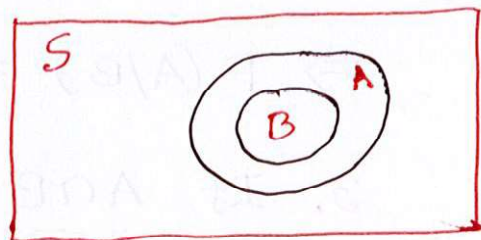
So, conditional probability

describes the probability of occurrence of A , given that the event B has already occurred

That is, Given (S, \mathcal{F}, P) and $A, B \in \mathcal{F}$ knowing that B has already occurred can tell us about whether the event A occurs or not.

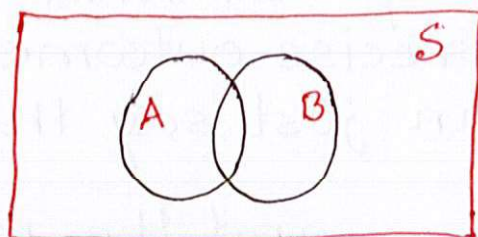


Here, $A \cap B = \emptyset$. So, if B occurs, then A does not.



Here, $A \cap B = B$, $B \subset A$. So, if B occurred, then A must have occurred.

Overall, knowledge on whether B has occurred or not may change our belief on the probable occurrence of A.



Definition of $P(A/B)$:

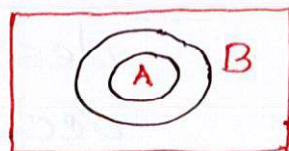
Given the probability space (S, \mathcal{F}, P) and the events $A, B \in \mathcal{F}$, the conditional probability of A conditioned on "B has occurred" is defined as:

$$P(A/B) = \frac{P(A \cap B)}{P(B)}, \quad P(B) \neq 0$$

↳ Probability of A given B

Cases: 1. If $A \subset B$, then $A \cap B = A$, so,

$$P(A/B) = \frac{P(A \cap B)}{P(B)}$$



$$\Rightarrow P(A/B) = \frac{P(A)}{P(B)} \gg P(A)$$

2. If $B \subset A$, then $A \cap B = B$, so

$$P(A/B) = \frac{P(A \cap B)}{P(B)}$$



$$\Rightarrow P(A/B) = \frac{P(B)}{P(B)}$$

$$\Rightarrow P(A/B) = 1$$

3. If $A \cap B = \emptyset$, then

$$P(A/B) = \frac{P(A \cap B)}{P(B)} = \frac{P(\emptyset)}{P(B)} = \frac{0}{P(B)} = 0$$

↳ Null is an impossible event

4. From the axioms of probability we know that $0 \leq P(B) \leq 1$

So, we can write for $P(A/B)$ definition that

$$0 < P(B) \leq 1$$

From definition of $P(A/B)$ we impose the constraint that $P(B) > 0$, as the event B must have occurred.

$$\text{Thus, } P(A/B) = \frac{P(A \cap B)}{P(B)} \geq P(A \cap B) \text{ as, } 0 < P(B) \leq 1$$

□ Axiomatic properties of $P(A/B)$

• Axiom 1: $P(A/B) \geq 0 \Rightarrow \frac{P(A \cap B)}{P(B)} \geq 0$

As we know, $P(A \cap B) \geq 0$ and $P(B) > 0$,

then, $P(A/B)$ should be ≥ 0

• Axiom 2: $P(S/B) = \frac{P(S \cap B)}{P(B)} = \frac{P(B)}{P(B)} = 1$

where, $B \subset S$

• Axiom 3: If $A = A_1 \cup A_2 \dots$ with $A_i \cap A_j = \emptyset$ for $i \neq j$, then

$$P(A/B) = P(A_1/B) + P(A_2/B) + \dots$$

Let's show it for a case where $A = A_1 \cup A_2$ with $A_1 \cap A_2 = \emptyset$

As, A_1 and A_2 are mutually exclusive, then $A_1 \cap B$ and $A_2 \cap B$ will also be mutually exclusive.

$$\begin{aligned}
\text{So, } P(A/B) &= P((A_1 \cup A_2)/B) \quad \rightarrow \text{distributive law} \\
&= \frac{P((A_1 \cup A_2) \cap B)}{P(B)} = \frac{P((A_1 \cap B) \cup (A_2 \cap B))}{P(B)} \\
&= \frac{P(A_1 \cap B) + P(A_2 \cap B)}{P(B)} = \frac{P(A_1 \cap B)}{P(B)} + \frac{P(A_2 \cap B)}{P(B)} \\
&= P(A_1/B) + P(A_2/B)
\end{aligned}$$

We can apply mathematical induction to expand the formulation for any n .

So, we can say that

$$(S, \mathcal{F}, P) \xrightarrow[\text{Given}]{B \text{ occurred}} (S, \mathcal{F}, P(\cdot/B))$$

Example: Let's calculate the conditional probability $P(A/B)$ for rolling a die experiment. $A = \{2\}$, $B = \{2, 4, 6\}$

As, it is a fair-die, $P(A) = \frac{1}{6}$

$$P(B) = \frac{3}{6} = \frac{1}{2}$$

$$\begin{aligned}
\text{So, } P(A/B) &= \frac{P(A \cap B)}{P(B)} = \frac{P(A)}{P(B)} \quad \left| \begin{array}{l} \text{as,} \\ A \subset B \end{array} \right. \\
&= \frac{1}{6} \times \frac{2}{1} = \frac{1}{3}
\end{aligned}$$

Baye's Theorem and Total probability Law

Assume that (S, \mathcal{F}, P) is a probability space, and A, B are events that are in the event space \mathcal{F} . So, $A, B \in \mathcal{F}$.

$$P(A/B) = \frac{P(A \cap B)}{P(B)} \quad \left| \quad P(B/A) = \frac{P(B \cap A)}{P(A)}\right.$$
$$\Rightarrow P(A \cap B) = P(B) P(A/B) \quad \left| \quad \begin{array}{l} \Rightarrow P(B \cap A) = P(A) P(B/A) \\ \Rightarrow P(A \cap B) = P(A) P(B/A) \end{array} \right.$$

Applied commutative law

We can now write

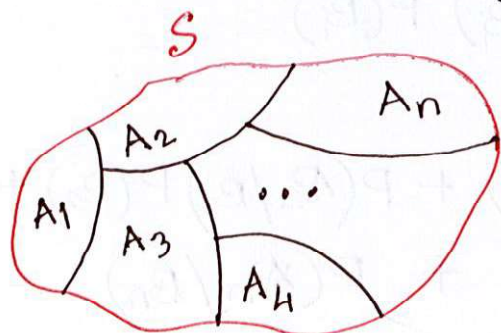
$$P(B) P(A/B) = P(A) P(B/A)$$
$$\Rightarrow P(A/B) = \frac{P(A) P(B/A)}{P(B)}$$

Also known as
Bayes Formula

Total Probability Law:

For (S, \mathcal{F}, P) , let's assume that $\{A_1, A_2, \dots, A_n\}$ be a partition of S , and the event $B \in \mathcal{F}(S)$, then

$$P(B) = P(B/A_1) P(A_1) + P(B/A_2) P(A_2) + \dots + P(B/A_n) P(A_n)$$



Proof of Total Probability Law

As B is an event and $B \in \mathcal{F}(S)$, we can say that $B \subset S$. So, $B = B \cap S$

$$\text{Then, } P(B) = P(B \cap S)$$

$$\Rightarrow P(B) = P\left(B \cap \left(\bigcup_{i=1}^n A_i\right)\right)$$

\hookrightarrow As $\{A_1, A_2, \dots, A_n\}$ is a partition

$$\Rightarrow P(B) = P\left(B \cap (A_1 \cup A_2 \dots \cup A_n)\right)$$

$$\Rightarrow P(B) = P\left((B \cap A_1) \cup (B \cap A_2) \cup \dots \cup (B \cap A_n)\right)$$

$$= P\left(\bigcup_{i=1}^n (B \cap A_i)\right)$$

A_1, A_2, \dots are mutually exclusive.

$$= \sum_{i=1}^n P(B \cap A_i) \dots \dots \textcircled{1}$$

So, $(B \cap A_1), (B \cap A_2), \dots$ will be mutually exclusive as well

From the definition of conditional probability

$$P(B / A_i) = \frac{P(A_i \cap B)}{P(A_i)} = \frac{P(B \cap A_i)}{P(A_i)}$$

$$\Rightarrow P(B \cap A_i) = P(B / A_i) P(A_i)$$

Using this expression in Eq.1, we obtain

$$P(B) = \sum_{i=1}^n P(B / A_i) P(A_i)$$

$$= P(B / A_1) P(A_1) + P(B / A_2) P(A_2) + \dots \\ \dots + P(B / A_n)$$

Example: Consider box 1 contains "a" white balls and "b" black balls.

Consider box 2 contains "c" white balls and "d" black balls.

Assume that one ball of unknown color is transferred from box 1 to box 2, and then,

Q. one ball is drawn from box 2. What is the probability that the drawn ball from box 2 would be white ball?

Answer: Let's define $W = \{ \text{transferred ball is white} \}$ → From box 1

$$\text{So, } P(W) = \frac{\text{\# of white balls}}{\text{\# of total balls}} = \frac{a}{a+b}$$

$$B = \{ \text{transferred ball is black} \} \Rightarrow \frac{b}{a+b} = P(B)$$

Now, sample space for box 1 $\equiv \{W, B\}$

Then, $W \cup B = S$, where W, B are disjoint

So, W, B forms a partition.

Consider that the desired event is

$$A = \{ \text{White ball is drawn from Box 2} \}$$

What information is known to us?

Transferred ball from box 1
(unknown color

known

$P(W)$

$P(B)$

W, B forms partition

From Bayes Formula, we know that

$$P(A/B) = \frac{P(B/A) P(A)}{P(B)}$$

So, for $A = A_i$,

$$P(A_i/B) = \frac{P(B/A_i) P(A_i)}{P(B)} \dots \dots \textcircled{1}$$

Now, using the total probability law we can write for a partition $\{A_1, A_2, \dots, A_n\}$ of S

$$P(B) = P(B/A_1) P(A_1) + P(B/A_2) P(A_2) + \dots$$

Thus,

$$P(A_i/B) = \frac{P(B/A_i) P(A_i)}{P(B/A_1) P(A_1) + P(B/A_2) P(A_2) \dots \dots}$$

$$= \frac{P(B/A_i) P(A_i)}{\sum_{i=1}^n P(B/A_i) P(A_i)} \quad \left| \begin{array}{l} \text{Bayes} \\ \text{Theorem} \end{array} \right.$$

So, Bayes Theorem is :

Let (S, \mathcal{F}, P) is a probability space and $\{A_1, A_2, \dots, A_n\}$ is a partition of the sample space S . Consider that $A_1, A_2, \dots, A_n \in \mathcal{F}$ and event B is in \mathcal{F} as well. So,

$$P(A_m/B) = \frac{P(B/A_m) P(A_m)}{\sum_{i=1}^n P(B/A_i) P(A_i)}$$

where, $m = 1, 2, \dots, n$

$$\begin{aligned}
 \text{So, } P(A) &= P(A \cap S) = P(A \cap (W \cup B)) \\
 &= P((A \cap W) \cup (A \cap B)) \quad \text{here, } A \cap W \text{ and } A \cap B \text{ are disjoint} \\
 &= P(A \cap W) + P(A \cap B)
 \end{aligned}$$

Using Bayes Formula

$$\begin{aligned}
 P(A \cap W) &= P(A/W) P(W) \\
 P(A \cap B) &= P(A/B) P(B)
 \end{aligned}$$

we can write,

$$P(A) = P(A/B) P(B) + P(A/W) P(W)$$

Now, $P(A/W)$ = Probability that white ball is drawn from box 2, given that the transferred ball from box 1 was white

$$= \frac{c+1}{c+d+1}$$

Similarly,

$$P(A/B) = \frac{c}{c+d+1}$$

Then,

$$P(A) = \frac{(c+1)}{(c+d+1)} \frac{a}{(a+b)} + \frac{c}{(c+d+1)} \frac{b}{(a+b)}$$

$$= \frac{a(c+1) + bc}{(c+d+1)(a+b)}$$

$$= \frac{ac + a + bc}{(c+d+1)(a+b)}$$

Answer.

Problem: Suppose, there are three types of players in a tournament.

A, B, C
A is 50% | c is 25%
B is 25%

Your winning chance against A, B, c types are 0.3, 0.4, 0.5 respectively.

Assume that, You play a match, so,

- What is the probability that you win the match?
- What is the probability that you played against A-type player? *Given, You won*

Answer: $P(A) = 0.5$, $P(B) = 0.25$, $P(c) = 0.25$

Let's define $W \equiv$ You win in the match.

So, $P(W/A) = 0.3$
 $P(W/B) = 0.4$
 $P(W/c) = 0.5$

A, B, c are the players type and they are disjoint. Also, their union will give all the player types

Then,

$$\begin{aligned} P(W) &= P(W/A)P(A) + P(W/B)P(B) + P(W/c)P(c) \\ &= (0.3 \times 0.5) + (0.4 \times 0.25) \\ &\quad + (0.5 \times 0.25) \\ &= 0.375 \end{aligned}$$

Partition
Concept

Total probability

Law

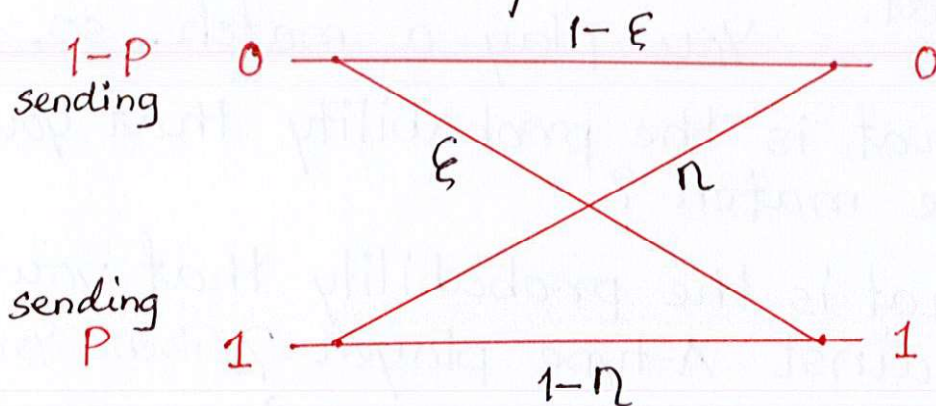
Answer

We have to find out $P(A/W) = ?$

Using Bayes Formula

$$P(A/W) = \frac{P(W/A) P(A)}{P(W)} = \frac{0.3 \times 0.5}{0.375}$$

Q. Consider a binary communication channel as:



As we see, Probability of sending 1 is p
Probability of sending 0 is $q = 1 - p$

Given that 1 is sent,
Probability of receiving 1 is $1 - \eta$

Given that 0 is sent
Probability of receiving 0 is $1 - \epsilon$

Probability of receiving 0 is $1 - \epsilon$

Probability of receiving 1 is ϵ

So, what is the probability that 1 has been correctly received?

Answer: We can not consider that $1 - \eta$ as the sole probability.

Because, 1 can be received because of a 0 transmission as well.

As transmitting either zero or one is mutually exclusive, we can use the concept of partition

So, $S_0 = 0$ is sent | $R_0 = 0$ is received
 $S_1 = 1$ is sent | $R_1 = 1$ is received

Using total probability law

$$\begin{aligned} P(R_1) &= P(R_1/S_1) P(S_1) + P(R_1/S_0) P(S_0) \\ &= (1-\eta)P + \epsilon(1-P) \end{aligned}$$

Let's extend it further

suppose we have received 1.

What is the probability that 1 was originally sent?

That is, we are interested in,

$$\begin{aligned} P(S_1/R_1) &= \frac{P(R_1/S_1) P(S_1)}{P(R_1)} \\ &= \frac{(1-\eta)P}{(1-\eta)P + \epsilon(1-P)} \end{aligned}$$

Q. Simplify the countable collection of unions

$$\begin{aligned}
 & \bigcup_{n=1}^{\infty} \left[5, 8 - (2n)^{-1} \right] = \bigcup_{n=1}^{\infty} \left[5, 8 - \frac{1}{2n} \right] \\
 & = \underbrace{\left[5, 8 - \frac{1}{2} \right]}_A \cup \underbrace{\left[5, 8 - \frac{1}{4} \right]}_C \cup \underbrace{\left[5, 8 - \frac{1}{6} \right]}_B \cup \dots \dots \\
 & = \left[5, 8 \right] \quad \text{As, } n \rightarrow \infty, \frac{1}{2 \times n} \cong 0 \\
 & \quad \quad \quad \text{So, } 8 - 0 = 8
 \end{aligned}$$

Q. Suppose, You import devices from three different sources:

	source	defective
Proportion of devices you import from different sources	A	0.005
	B	0.001
	C	0.01

$$\begin{array}{l|l}
 A : 0.5 & C : 0.4 \\
 B : 0.1 &
 \end{array}$$

a) Probability that a randomly selected device is defective and that it is from source A is: $P(D/A) = ?$

$$\begin{aligned}
 P(D) &= P(D/A)P(A) + P(D/B)P(B) + P(D/C)P(C) \\
 &= (0.005)0.5 + (0.001)0.1 + (0.01)0.4 \\
 &= 0.0066
 \end{aligned}$$

$$P(A/D) = \frac{P(D/A)P(A)}{P(D)} = \frac{0.005 \times 0.5}{0.0066}$$

=

$$\begin{aligned} \text{b) } P(B/D) &= \frac{P(D/B) P(B)}{P(D)} \\ &= \frac{0.001 \times 0.0152}{0.0066} = 0.0152 \end{aligned}$$

$$\text{c) } P(C/D) = \frac{P(D/C) P(C)}{P(D)} = \frac{0.01 \times 0.4}{0.0066} = 0.6$$